

Theoretical Statistics, STOR 655

Total Variation Distance and Kullback-Liebler Divergence

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Background

Basic question: How far apart (different) are two distributions P and Q ?

- ▶ Measured through distances and divergences
- ▶ Used to define convergence of distributions
- ▶ Used to assess smoothness of parametrizations $\{P_\theta : \theta \in \Theta\}$
- ▶ Means of assessing the complexity of a family of distributions
- ▶ Key ingredient in formulating lower and upper bounds on the performance of inference procedures

Kolmogorov-Smirnov Distance

Definition: Let P and Q be probability distributions on \mathbb{R} with CDFs F and G . The Kolmogorov-Smirnov (KS) distance between P and Q is

$$\text{KS}(P, Q) = \sup_t |F(t) - G(t)|$$

Properties of KS distance

1. $0 \leq \text{KS}(P, Q) \leq 1$
2. $\text{KS}(P, Q) = 0$ iff $P = Q$
3. KS is a metric
4. $\text{KS}(P, Q) = 1$ iff exists $s \in \mathbb{R}$ with $P((-\infty, s]) = 1$ and $Q((s, \infty)) = 1$

Total Variation Distance

Definition: Let \mathcal{X} be a set with a sigma-field \mathcal{A} . The total variation distance between two probability measures P and Q on $(\mathcal{X}, \mathcal{A})$ is

$$\text{TV}(P, Q) = \sup_{A \in \mathcal{A}} |P(A) - Q(A)|$$

Properties of Total Variation

1. $0 \leq \text{TV}(P, Q) \leq 1$
2. $\text{TV}(P, Q) = 0$ iff $P = Q$
3. TV is a metric
4. $\text{TV}(P, Q) = 1$ iff there exists $A \in \mathcal{A}$ with $P(A) = 1$ and $Q(A) = 0$

KS, TV, and the CLT

Note: $KS(P, Q)$ and $TV(P, Q)$ can both be expressed in the form

$$\sup_{A \in \mathcal{A}_0} |P(A) - Q(A)|$$

For KS sup over all intervals $(-\infty, t]$, while for TV sup over all Borel sets

Example: Let $X_1, X_2, \dots \in \{-1, 1\}$ iid with $\mathbb{P}(X_i = 1) = \mathbb{P}(X_i = -1) = 1/2$.
By the standard central limit theorem

$$Z_n = \frac{1}{n^{1/2}} \sum_{i=1}^n X_i \Rightarrow \mathcal{N}(0, 1)$$

Let $P_n =$ distribution of Z_n and $Q = \mathcal{N}(0, 1)$. Can show that

$$KS(P_n, Q) \leq cn^{-1/2} \quad \text{while} \quad TV(P_n, Q) \equiv 1$$

Total Variation and Densities

Scheffé's Theorem: Let $P \sim f$ and $Q \sim g$ be distributions on $\mathcal{X} = \mathbb{R}^d$. Then

1. $\text{TV}(P, Q) = \frac{1}{2} \int |f(x) - g(x)| dx$

2. $\text{TV}(P, Q) = 1 - \int \min\{f(x), g(x)\} dx$

3. $\text{TV}(P, Q) = P(A) - Q(A)$ where $A = \{x : f(x) \geq g(x)\}$

Analogous results hold when $P \sim p(x)$ and $Q \sim q(x)$ are described by pmfs

Upshot: Total variation distance between P and Q is half the L_1 -distance between densities or mass functions

Total Variation and Hypothesis Testing

Problem: Observe $X \in \mathcal{X}$ having density f_0 or f_1 . Wish to test

$$H_0 : X \sim f_0 \text{ vs. } H_1 : X \sim f_1$$

Any decision rule $d : \mathcal{X} \rightarrow \{0, 1\}$ has overall (Type I + Type II) error

$$\text{Err}(d) = \mathbb{P}_0(d(X) = 1) + \mathbb{P}_1(d(X) = 0)$$

Fact: The optimum overall error among *all* decision rules is

$$\inf_{d: \mathcal{X} \rightarrow \{0,1\}} \text{Err}(d) = \int \min\{f_0(x), f_1(x)\} dx = 1 - \text{TV}(P_0, P_1)$$

Coupling and Total Variation

Fact: Let P and Q be distributions on \mathcal{X} . Then

$$\text{TV}(P, Q) = \min_{(X, Y)} \mathbb{P}(X \neq Y)$$

where the minimum is over all joint distributions (X, Y) such that $X \sim P$ and $Y \sim Q$. A joint distribution of this sort is called a *coupling*

Corollary

- ▶ If $X \sim P$ and $Y \sim Q$ are defined on the same probability space then $\mathbb{P}(X = Y) \leq 1 - \text{TV}(P, Q)$
- ▶ There is an optimal coupling achieving the upper bound, which makes X and Y equal as much as possible

Kullback-Liebler (KL) Divergence

Definition: The *KL-divergence* between distributions $P \sim f$ and $Q \sim g$ is

$$\text{KL}(P : Q) = \int f(x) \log \frac{f(x)}{g(x)} dx = \mathbb{E}_f \left[\log \frac{f(X)}{g(X)} \right]$$

Analogous definition for discrete distributions $P \sim p$ and $Q \sim q$

- ▶ The integrand can be positive or negative. By convention

$$f(x) \log \frac{f(x)}{g(x)} = \begin{cases} +\infty & \text{if } f(x) > 0 \text{ and } g(x) = 0 \\ 0 & \text{if } f(x) = 0 \end{cases}$$

- ▶ KL divergence is not symmetric, and is not a metric

First Properties of KL Divergence

Fact: Divergence $\text{KL}(P : Q)$ is well defined: if $u_- = \max(-u, 0)$ then

$$\int \left(f(x) \log \frac{f(x)}{g(x)} \right)_- dx \leq 1$$

Key Fact:

- ▶ Divergence $\text{KL}(P : Q) \geq 0$ with equality if and only if $P = Q$
- ▶ $\text{KL}(P : Q) = +\infty$ if there is a set A with $P(A) > 0$ and $Q(A) = 0$

Notation: When pmfs/pdfs clear from context, write $\text{KL}(p : q)$ or $\text{KL}(f : g)$

KL Divergence Examples

Example: Let p and q be pmfs on $\{0, 1\}$ with

$$p(0) = p(1) = 1/2 \quad \text{and} \quad q(0) = (1 - \epsilon)/2, \quad q(1) = (1 + \epsilon)/2$$

where $\epsilon \in (0, 1)$. Then we have

▶ $\text{KL}(p : q) = -\frac{1}{2} \log(1 - \epsilon^2) \leq \epsilon^2$ when $\epsilon \leq \frac{1}{\sqrt{2}}$

▶ $\text{KL}(q : p) = \frac{1}{2} \log(1 - \epsilon^2) + \frac{\epsilon}{2} \log\left(\frac{1-\epsilon}{1+\epsilon}\right) \leq 2\epsilon^2$

Example: If $P \sim \mathcal{N}_d(\mu_0, \Sigma_0)$ and $Q \sim \mathcal{N}_d(\mu_1, \Sigma_1)$ with $\Sigma_0, \Sigma_1 > 0$ then

$$2 \text{KL}(P : Q) = \text{tr}(\Sigma_1^{-1} \Sigma_0) + (\mu_1 - \mu_0)^t \Sigma_1^{-1} (\mu_1 - \mu_0) + \ln(|\Sigma_1|/|\Sigma_0|) - d$$

KL Divergence and Inference

Ex 1. (Testing) Consider testing $H_0 : X \sim f_0$ vs. $H_1 : X \sim f_1$. The divergence

$$\text{KL}(f_0 : f_1) = \mathbb{E}_0 \left(\log \frac{f_0(X)}{f_1(X)} \right) \geq 0$$

is just the expected log likelihood ratio under H_0

Ex 2. (Estimation) Let X_1, X_2, \dots iid with $X_i \sim f(x|\theta_0) \in \{f(x|\theta) : \theta \in \Theta\}$. Under suitable assumptions, when n is large,

$$\hat{\theta}_n^{\text{MLE}}(X_1^n) \approx \underset{\theta \in \Theta}{\text{argmin}} \text{KL}(f(\cdot|\theta_0) : f(\cdot|\theta))$$

In other words, MLE is trying to find θ minimizing KL divergence with true distribution

Data Processing Inequality

- ▶ Measurable spaces $(\mathcal{X}, \mathcal{A})$ with measures P and Q
- ▶ Measurable function $f : \mathcal{X} \rightarrow \mathcal{Y}$ from $(\mathcal{X}, \mathcal{A})$ to $(\mathcal{Y}, \mathcal{B})$
- ▶ Map f pushes P and Q forward to measures \tilde{P} and \tilde{Q} on $(\mathcal{Y}, \mathcal{B})$ where

$$\tilde{P}(B) = P(f^{-1}B) \quad \text{and} \quad \tilde{Q}(B) = Q(f^{-1}B)$$

Data Processing Inequality: Application of f reduces divergence, namely

$$\text{KL}(\tilde{P} : \tilde{Q}) \leq \text{KL}(P : Q)$$

Result extends to stochastic transformations (transition kernels) from \mathcal{X} to \mathcal{Y}

Variational Formulation and Convexity

Fact: Let P and Q be distributions on $(\mathcal{X}, \mathcal{A})$. Then

$$\text{KL}(P : Q) = \sup_f \left[\int f dP - \log \left(\int e^f dQ \right) \right]$$

where the supremum is over all functions $f : \mathcal{X} \rightarrow \mathbb{R}$ such that $\int e^f dQ$ is finite

Corollary: For each distribution Q on $(\mathcal{X}, \mathcal{A})$ the function $\text{KL}(\cdot : Q)$ is convex: if P_1, P_2 are distributions and $\alpha \in (0, 1)$ then

$$\text{KL}(\alpha P_1 + (1 - \alpha) P_2 : Q) \leq \alpha \text{KL}(P_1 : Q) + (1 - \alpha) \text{KL}(P_2 : Q)$$

Product Densities (Tensorization)

Notation: Given distributions P_1, \dots, P_n on \mathcal{X} with densities f_1, \dots, f_n let $\otimes_{i=1}^n P_i$ denote the product distribution on \mathcal{X}^n with density $f_1(x_1) \cdots f_n(x_n)$

Tensorization: Let P_1, \dots, P_n and Q_1, \dots, Q_n be distributions on \mathcal{X} with densities f_1, \dots, f_n and g_1, \dots, g_n , respectively. Then

1. $\text{KS}(\otimes_{i=1}^n P_i, \otimes_{i=1}^n Q_i) \leq \sum_{i=1}^n \text{KS}(P_i, Q_i)$
2. $\text{TV}(\otimes_{i=1}^n P_i, \otimes_{i=1}^n Q_i) \leq \sum_{i=1}^n \text{TV}(P_i, Q_i)$
3. $\text{KL}(\otimes_{i=1}^n P_i : \otimes_{i=1}^n Q_i) = \sum_{i=1}^n \text{KL}(P_i, Q_i)$

Kullback Liebler vs Total Variation

Pinsker's Inequality: For any distributions P and Q on $(\mathcal{X}, \mathcal{A})$,

$$\text{KL}(P : Q) \geq 2\text{TV}(P : Q)^2$$