Theoretical Statistics, STOR 655 Asymptotic Analysis of the  $T^2$  and  $\chi^2$  Statistics

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# Limiting Distribution of Hotelling's $T^2$

#### Preliminaries

**Fact:** If  $X \sim \mathcal{N}_d(\mu, \Sigma)$  with  $\Sigma > 0$  then

$$W = (X - \mu)\Sigma^{-1}(X - \mu)^{t} \sim \chi_{d}^{2}$$

**Definition:** The sample covariance matrix of  $X_1, \ldots, X_n \in \mathbb{R}^d$  is

$$S_{n} = \frac{1}{n} \sum_{i=1}^{n} (X_{i} - \overline{X}_{n}) (X_{i} - \overline{X}_{n})^{t} = \frac{1}{n} \sum_{i=1}^{n} X_{i} X_{i}^{t} - (\overline{X}_{n}) (\overline{X}_{n})^{t}$$

**Fact:** If  $X_1, X_2, \ldots \in \mathbb{R}^d$  are iid with common variance matrix  $Var(X_i) = \Sigma$  then  $S_n \to \Sigma$  wp1 as  $n \to \infty$ 

## Hotelling's $T^2$

**Definition:** Let  $X_1, \ldots, X_n \in \mathbb{R}^d$  be iid, and let  $\mu \in \mathbb{R}^d$  be a mean vector of interest. Hotelling's  $T^2$  statistic is

$$T_n^2 = (n-1)(\overline{X}_n - \mu)^t S_n^{-1}(\overline{X}_n - \mu)$$

where  $S_n$  is the sample variance matrix based on  $X_1, \ldots, X_n$ 

- Multivariate analog of the one-sample t-statistic
- Used for inference about the common mean of the X<sub>i</sub>'s

Fact: If  $X_1, X_2, \ldots \in \mathbb{R}^d$  are iid with  $\mathbb{E}X = \mu$  and  $\operatorname{Var}(X_i) = \Sigma > 0$  then  $T_n^2 \Rightarrow \chi_d^2$ 

Limiting Distribution of Pearson's  $\chi^2$ 

### **Review: Projection Matrices**

**Definition:**  $A \in \mathbb{R}^{d \times d}$  is a projection matrix if  $A^2 = A$ 

**Idea:** Suppose  $A^2 = A$ , and let V = span of columns of A

• Matrix A maps vector  $u \in \mathbb{R}^d$  to vector  $v = Au \in V$ 

• Matrix A leaves vectors in V unchanged: if  $v \in V$  then v = Au so

$$Av = A(Au) = A^2u = Au = v$$

• Matrix A projects  $\mathbb{R}^d$  onto subspace V



Fact: Let A be a projection matrix

- 1. All eigenvalues of A are 0 or 1
- 2. rank(A) = trace(A)
- 3. If A is symmetric then  $Ax \perp (x Ax)$  for every  $x \in \mathbb{R}^d$

**Fact:** Let  $X \sim \mathcal{N}_d(0, \Sigma)$ . Then  $X^t X \sim \chi_r^2$  iff  $\Sigma$  is a projection of rank r

### Multinomial Experiment

#### **Multinomial Experiment**

Sequence of n iid trials where each trial has one of d possible outcomes

- Let  $p_k$  = probability of outcome k on any given trial
- Let  $p = (p_1, \ldots, p_d)^t$  be pmf of trial outcomes
- Let  $n_k$  = number of trials having outcome k. Thus  $\sum_{k=1}^d n_k = n$

**Definition:** Multinomial(n, p) is the joint distribution of  $(n_1, \ldots, n_d)$ 

$$P(x_1,...,x_d) = \frac{n!}{x_1!\cdots x_d!} p_1^{x_1}\cdots p_d^{x_d}$$

## Multinomial Goodness of Fit via the $\chi^2$ Statistic

#### Inference problem

- Perform multinomial experiment. Observe counts n<sub>1</sub>,..., n<sub>d</sub>
- Assess fit of n<sub>1</sub>,..., n<sub>d</sub> to a Multinomial(n, p) distribution, where p is a fixed pmf of interest

Note: Under the Multinomial(n, p) distribution,  $\mathbb{E}(n_k) = n p_k$ 

**Definition:** The  $\chi^2$  statistic is given by

$$\chi_n^2(n_1,\ldots,n_d) = \sum_{k=1}^d \frac{(n_k - np_k)^2}{np_k} = \sum \frac{(\mathsf{observed} - \mathsf{expected})^2}{\mathsf{expected}}$$

### Limiting Distribution of $\chi^2$

**Theorem:** If  $n_1, \ldots, n_d$  are obtained from the target Multinomial(n, p) distribution then  $\chi_n^2 \Rightarrow \chi_{d-1}^2$  as the number of trials *n* tends to infinity

**Modified**  $\chi^2$ . Let  $g : \mathbb{R}^d \to \mathbb{R}^d$  be of the form  $g(x) = (g_1(x_1), \dots, g_d(x_d))^t$ where  $g_k : \mathbb{R} \to \mathbb{R}$ . Using the delta method, one can show

$$\chi_n^2(g) = n \sum_{k=1}^d \frac{(g_k(n_k/n) - g_k(p_k))^2}{p_k g'_k(p_k)^2} \Rightarrow \chi_{d-1}^2$$

Special case  $g_k(x) = x^{1/2}$  for  $1 \le k \le d$  gives Hellinger's  $\chi^2$ 

$$\chi^2(g) = 4n \sum_{k=1}^d (\sqrt{n_k/n} - \sqrt{p_k})^2$$