

## STOR 654 Homework 10

1. *Extreme value theory for the Gaussian.* Let  $a_n$  and  $b_n$  be the extreme value scaling and centering constants for the maximum  $M_n$  of  $n$  independent standard Gaussian random variables.

(a) Fix  $x \in \mathbb{R}$  and let  $x_n = x/a_n + b_n$ . Show that  $n\phi(x_n)/x_n \rightarrow e^{-x}$  as  $n$  tends to infinity. [In your calculations, identify and pay careful attention to the leading order terms.]

(b) Using the result of part (a) and the standard Gaussian tail bound from an earlier homework, show that  $n(1 - \Phi(x_n)) \rightarrow e^{-x}$ .

(c) Use part (b) and the lemma from lecture to show that as  $n$  tends to infinity

$$\mathbb{P}(a_n(M_n - b_n) \leq x) \rightarrow G(x) = e^{-e^{-x}}$$

(d) Show that  $G(x)$  is the CDF of  $-\log V$  where  $V \sim \text{Exp}(1)$ .

2. Let  $M_n$  be the maximum of  $n$  iid  $\mathcal{N}(0, 1)$  random variables. Use the Gaussian extreme value theorem to establish the following limiting results.

a.  $\mathbb{P}(M_n \geq \sqrt{2 \log n}) \rightarrow 0$  as  $n \rightarrow \infty$

b.  $M_n/\sqrt{2 \log n} \rightarrow 1$  in probability as  $n \rightarrow \infty$

3. Let  $V \subseteq \mathbb{R}^n$  be a finite set of vectors  $v = (v_1, \dots, v_n)^t$  with  $L = \max_{v \in V} \|v\|_2$ , and let  $\varepsilon_1, \dots, \varepsilon_n$  be independent Rademacher (sign) variables.

(a) Use Hoeffding's MGF inequality to bound the moment generating functions of the random variables  $\sum_{i=1}^n \varepsilon_i v_i$  in terms of the constant  $L$ .

(b) Show that

$$\mathbb{E} \left[ \max_{v \in V} \sum_{i=1}^n \varepsilon_i v_i \right] \leq \sqrt{2L^2 \log |V|}$$

4. Let  $U_{(r)}$  be the  $r$ 'th order statistics of  $U_1, \dots, U_n$  iid  $\sim U(0, 1)$ . Show that  $\mathbb{E}U_{(r)} = r/(n+1)$ . (A formula for the density of  $U_{(r)}$  was given in an earlier lecture.)

5. Define the function  $g_s(z) = \exp(sz - s^2/2)$ , where  $s$  is a constant, and let  $Z \sim \mathcal{N}(0, 1)$ . Show that for  $t \in \mathbb{R}$

- a.  $\mathbb{E}g_s(Z)\mathbb{I}(Z \leq t) = \Phi(\epsilon t)$
- b.  $\mathbb{E}g_s^2(Z)\mathbb{I}(Z \leq t) = e^{s^2}\Phi(t - 2s)$
- c.  $\Phi(t - 2s) \leq \phi(2s - t)$  when  $t - 2s \leq -1$

Fix  $\epsilon \in (0, 1/2)$ . For  $n \geq 1$  let  $t_n = \sqrt{2 \log n}$ ,  $s_n = (1 - \epsilon)t_n$ , and define

$$\tilde{L}_n = n^{-1} \sum_{i=1}^n g_{s_n}(Z_i) \mathbb{I}(Z_i \leq t_n)$$

where  $Z_1, \dots, Z_n$  iid  $\sim \mathcal{N}(0, 1)$ . Using (a)-(c) above as needed, establish the following:

- d.  $\mathbb{E}\tilde{L}_n = \Phi(\epsilon t) \rightarrow 1$  as  $n \rightarrow \infty$
- e.  $\text{Var}(\tilde{L}_n) \leq (2\pi)^{-1}e^{-\epsilon^2 t^2} \rightarrow 0$  as  $n \rightarrow \infty$
- f.  $\tilde{L}_n \rightarrow 1$  in probability as  $n \rightarrow \infty$

6. Let  $x < y$  be real numbers. Carefully justify each of the following (in)equalities.

$$\frac{e^y - e^x}{y - x} = \int_0^1 e^{tx + (1-t)y} dt \leq \int_0^1 (te^x + (1-t)e^y) dt = \frac{1}{2}(e^x + e^y)$$

7. Let  $X$  and  $Y$  be random variables, possibly defined on different probability spaces, with CDFs  $F$  and  $G$ , respectively. We say that  $X$  is stochastically larger than  $Y$ , written  $X \stackrel{d}{\geq} Y$  if  $F(x) \leq G(x)$  for each  $x \in \mathbb{R}$ . Explain the intuition behind the definition. Recall that if  $X$  is a random variable with continuous CDF  $F$  then  $F(X) \stackrel{d}{=} U(0, 1)$ . In general, even if  $F$  is not continuous, it is still true that  $F(X) \stackrel{d}{\geq} U(0, 1)$ . Establish this fact. Hint: Let  $\varphi$  be the percentile function of  $F$ . Argue that  $F(\varphi(u)) \geq u$  for  $0 < u < 1$ , and recall that  $X \stackrel{d}{=} \varphi(U)$  where  $U \sim U(0, 1)$ .

8. *Independent Copies.* Let  $X, X'$  be independent random variables with the same distribution. In this case we say that  $X'$  is an independent copy of  $X$ .

- (a) Show that  $\text{Var}(X) = \frac{1}{2}\mathbb{E}(X - X')^2$
- (b) Argue formally or informally that  $\mathbb{E}(X' | X) = \mathbb{E}X$
- (c) Using the result of part (b) and Jensen's inequality for conditional expectations, show that  $\mathbb{E}|X - \mathbb{E}X| \leq \mathbb{E}|X - X'|$ . This is a key step in establishing a number of important bounds in empirical process theory.

9. (Using symmetry) Let  $X_1, \dots, X_n$  be iid positive random variables such that  $\mathbb{E}X_k = a$  and  $\mathbb{E}X_k^{-1} = b$  are finite. For  $1 \leq k \leq n$ , let  $S_k = X_1 + \dots + X_k$  and let  $V_k = X_k/S_n$ .

- (a) Show that  $\mathbb{E}S_m^{-1} < \infty$  for  $m = 1, \dots, n$ .
- (b) Argue informally that  $V_1, \dots, V_n$  have the same distribution.
- (c) Conclude from part (b) that  $\mathbb{E}(S_m/S_n) = m/n$  for  $1 \leq m \leq n$ .
- (d) Show that  $\mathbb{E}(S_n/S_m) = 1 + (n - m)a\mathbb{E}S_m^{-1}$  for  $1 \leq m \leq n$ .
- (e) Verify the inequality  $x + x^{-1} \geq 2$  for  $x > 0$ .
- (f) Use the inequality of part (e) to show that  $\mathbb{E}(S_n/S_m) \geq n/m$  for  $1 \leq m \leq n$ .

10. Let  $U \sim U(0, 1)$ . Show that  $\max(U, 1 - U) \stackrel{d}{=} 1/2 + U/2$ .